

Sharper rates for estimating differential entropy under Gaussian convolutions

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Abstract

In this short note, we show that, given access to n i.i.d. samples from a compactly supported d -dimensional distribution P , the differential entropy of P convolved with an isotropic Gaussian can be estimated at the rate $O(n^{1/2})$ by a plug-in estimator. This answers a question of Goldfeld et al. (2018).

We consider the following problem: given i.i.d. samples from a distribution P on $[-1, 1]^d$, how well can one estimate the differential entropy of P convolved with an isotropic Gaussian? If we denote by \mathcal{N}_σ the distribution $\mathcal{N}(0, \sigma^2 I_d)$ and by $*$ the convolution operator, Goldfeld et al. (2018) recently showed that there exists a simple estimator which converges to $h(P * \mathcal{N}_\sigma)$ at nearly the parametric rate. Indeed, writing $P_n := \frac{1}{n} \sum_{i=1}^n \delta_{X_i}$, where $X_i \sim P$ i.i.d., they showed (Goldfeld et al., 2018, Theorem 2) that the *plug-in estimator* $h(P_n * \mathcal{N}_\sigma)$ achieves:

$$\mathbb{E}|h(P_n * \mathcal{N}_\sigma) - h(P * \mathcal{N}_\sigma)| \leq c_{\sigma,d} \frac{(\log n)^{\frac{d}{4}}}{\sqrt{n}}. \quad (1)$$

This rate is striking in that it is significantly better than what could be achieved by a generic estimator using samples from $P * \mathcal{N}_\sigma$ alone (see, e.g., Han et al., 2017). In the interest of obtaining sharp rates, Goldfeld et al. (2018) posed the question of whether the logarithmic term $(\log n)^{\frac{d}{4}}$ could be improved to $(\log n)^c$ for some universal constant c .

In this note, we answer this question in the affirmative, showing in fact that the plug-in estimator $h(P_n * \mathcal{N}_\sigma)$ achieves exactly the parametric rate, without logarithmic factors.

Theorem 1. *For any distribution P supported on $[-1, 1]^d$, we have*

$$\mathbb{E}|h(P_n * \mathcal{N}_\sigma) - h(P * \mathcal{N}_\sigma)| \leq c_{\sigma,d} \frac{1}{\sqrt{n}},$$

for $c_{\sigma,d} := \frac{d \cdot 2^{d+3}}{\min\{\sigma^2, \sigma^{d+2}\}}$.

The proof of Theorem 1 relies on the following proposition. Denote by $W_1(P, Q)$ the 1-Wasserstein distance between P and Q , i.e., $W_1(P, Q) := \inf_\gamma \int \|x - y\| d\gamma(x, y)$, where the infimum is taken over all couplings of P and Q .

Proposition 1. *If P is supported on $[-1, 1]^d$, then*

$$\mathbb{E}W_1(P_n * \mathcal{N}_\sigma, P * \mathcal{N}_\sigma) \leq c'_{\sigma,d} \frac{1}{\sqrt{n}},$$

for $c'_{\sigma,d} := \frac{\sqrt{d} \cdot 2^{d+2}}{\min\{1, \sigma^d\}}$.

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To connect Proposition 1 to the question of entropy estimation, we employ the following result due to Polyanskiy and Wu (2016).

Proposition 2 (Polyanskiy and Wu, 2016, Proposition 5). *Let P and Q be distributions supported on $[-1, 1]^d$, with $v_P := \mathbb{E}_{X \sim P} \|X\|^2$ and $v_Q := \mathbb{E}_{X \sim Q} \|X\|^2$. Then*

$$|h(Q * \mathcal{N}_\sigma) - h(P * \mathcal{N}_\sigma)| \leq \frac{1}{2\sigma^2} (|v_Q - v_P| + 2\sqrt{d}W_1(Q * \mathcal{N}_\sigma, P * \mathcal{N}_\sigma)).$$

When $Q = P_n$, Jensen's inequality implies $\mathbb{E}|v_{P_n} - v_P| \leq \frac{1}{\sqrt{n}} \text{var}_{X \sim P} (\|X\|^2)^{1/2} \leq d/\sqrt{n}$. Hence, Theorem 1 follows directly from Propositions 1 and 2. It therefore suffices to give a proof of Proposition 1.

Proof of Proposition 1. Denote by f the density of $P * \mathcal{N}_\sigma$, and by f_n the density of $P_n * \mathcal{N}_\sigma$. We let $\phi_\sigma(x) := (2\pi\sigma^2)^{-d/2} \exp(-\frac{1}{2\sigma^2}\|x\|^2)$ be the density of \mathcal{N}_σ . We use the following upper bound (Villani, 2008, Theorem 6.15):

$$W_1(P_n * \mathcal{N}_\sigma, P * \mathcal{N}_\sigma) \leq \int_{\mathbb{R}^d} \|z\| |f_n(z) - f(z)| dz.$$

This yields

$$\begin{aligned} \mathbb{E}W_1(P_n * \mathcal{N}_\sigma, P * \mathcal{N}_\sigma) &\leq \int_{\mathbb{R}^d} \|z\| \mathbb{E}|f_n(z) - f(z)| dz \\ &= \int_{\mathbb{R}^d} \|z\| \mathbb{E} \left| \frac{1}{n} \sum_{i=1}^n \phi_\sigma(z - X_i) - \mathbb{E}\phi_\sigma(z - X) \right| dz \\ &\leq \frac{1}{\sqrt{n}} \int_{\mathbb{R}^d} \|z\| (\mathbb{E}(\phi_\sigma(z - X) - \mathbb{E}\phi_\sigma(z - X))^2)^{1/2} dz, \quad X \sim P \\ &\leq \frac{1}{\sqrt{n}} \int_{\mathbb{R}^d} \|z\| (\mathbb{E}\phi_\sigma(z - X)^2)^{1/2} dz. \end{aligned}$$

When $z \in [-2, 2]^d$, we use the bound $(\mathbb{E}\phi_\sigma(z - X)^2)^{1/2} \leq \max_{z \in \mathbb{R}^d} \phi_\sigma(z) = (2\pi\sigma^2)^{-d/2}$. For $z \notin [-2, 2]^d$, we have $\|z - X\|^2 \geq \|z/2\|^2$ almost surely, which yields $(\mathbb{E}\phi_\sigma(z - X)^2)^{1/2} \leq \phi_\sigma(z/2)$. We obtain

$$\begin{aligned} \mathbb{E}W_1(P_n * \mathcal{N}_\sigma, P * \mathcal{N}_\sigma) &\leq \frac{(2\pi\sigma^2)^{-d/2}}{\sqrt{n}} \int_{z \in [-2, 2]^d} \|z\| dz + \frac{1}{\sqrt{n}} \int_{z \in \mathbb{R}^d} \|z\| \phi_\sigma(z/2) dz \\ &\leq \left((2\pi\sigma^2)^{-d/2} \cdot 4^d \cdot 2 + 2^{d+1} \right) \cdot \sqrt{d/n} \\ &\leq \max\{1, \sigma^{-d}\} 2^{d+2} \sqrt{d/n}. \end{aligned}$$

□

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